## Hypothesis Tests - Types (8 pages; 24/1/20)

See also "Hypothesis Tests - General".

(excluding Correlation,  $\chi^2$  test for independence or Goodness of Fit, and Analysis of Variance)

[Unless indicated otherwise, a 'large' sample means  $n \ge 30$ ]

## (A) Single samples

(1) Test for mean of a Normal distribution with known variance

$$H_0: X \sim N(\mu, \sigma^2) \Rightarrow \overline{X} \sim N(\mu, \frac{\sigma^2}{n})$$

Test statistic:  $z = \frac{\overline{x} - \mu}{\left(\frac{\sigma}{\sqrt{n}}\right)}$  (compare with the critical value from the Normal table).

(2) Test for mean of a Normal distribution with unknown variance, where the sample is large

As the sample is large (usually taken to be  $\geq 30$ ), s (based on a divisor of n - 1) can be assumed to be a reasonably good approximation to  $\sigma$ .

The test statistic is then  $z = \frac{\overline{x} - \mu}{\left(\frac{s}{\sqrt{n}}\right)}$ 

(3) Test for mean of an unknown distribution (with unknown variance), where the sample is large

As the sample size is large, the Central Limit theorem says that  $\overline{X}$  approx.  $\sim N(\mu, \frac{\sigma^2}{n})$  [ $n \geq 30$  also applies here] and s can be assumed to be a reasonably good approximation to  $\sigma$  (again, as the sample size is large).

As in (2), the test statistic is  $z = \frac{\overline{x} - \mu}{\left(\frac{s}{\sqrt{n}}\right)}$ 

(4) Test for mean of a Normal distribution with unknown variance, where the sample is small

To reflect the greater uncertainly caused by approximating  $\sigma$  by s when the sample is small, the *t*-distribution is used, with

v = n - 1 degrees of freedom.

The test statistic is  $t_{n-1} = \frac{\overline{x} - \mu}{\left(\frac{s}{\sqrt{n}}\right)}$  (compare with the critical value from the *t* table).

(Note that the underlying distribution has to be Normal, in order for the *t*-distribution to apply. As the sample size increases, the

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t-value tends to the z-value.)
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(5) Test for mean or median of a symmetrical but otherwise unknown distribution, where the sample is small

The Wilcoxon signed-rank test can be used.

This is a test for the median, but can be used approximately for the mean.

Procedure:

(i)  $H_0$ : median is M

(ii) Given a sample of  $x_i$  of size n , calculate the differences  $x_i - M$ 

(iii) Rank the  $x_i - M$  by absolute size, with a rank of 1 for the smallest value of  $|x_i - M|$ 

(iv) Calculate the sum of ranks for the positive differences, and also for the negative differences. The test statistic, W is then the smaller of these sums.

(v) The (lower tail) critical value is obtained from the Wilcoxon signed-rank table. Reject  $H_0$  if W < critical value [ie W is suspiciously small if  $H_0$  is assumed]

[Note: For larger n,

 $W \sim \operatorname{approx.} N(\frac{1}{4}n(n+1), \frac{1}{24}n(n+1)(2n+1))$ , with a continuity correction being applied (presumably); however, (3) may also be applied for large n]

(6) Test for mean or median of an unknown distribution (which cannot be assumed to be symmetrical), where the sample is small

The Sign test can be used.

 $H_0$ : median is M

Test statistic is P: the number of positive values of  $x_i - M$ 

 $H_0 \Rightarrow P \sim B(n, 0.5),$ 

so that the critical value is obtained from the Binomial table

(7) Test for Binomial proportion, for a large sample (using a Normal approximation)

 $H_0: X \sim B(n, p)$ . If *n* is large and *p* is not too small, in such a way that a Normal approximation is appropriate (this will usually be the case if  $n \ge 50 \& np \ge 10$ ), then *X* approx.  $\sim N(np, np(1-p))$ .

[A continuity correction is not usually required.]

So the test statistic is X, the number of successes, and the critical value is obtained from the Normal table. [Note: For the earlier tests, a sample is required; but here we are using a single value resulting from *n* trials.]

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[Alternatively,  $Y = \frac{X}{n}$  can be taken to be the test statistic, and use made of the fact that Y approx.  $\sim N\left(p, \frac{p(1-p)}{n}\right)$ : this approach is used for determining a confidence interval for the population proportion (see "Confidence Intervals").]

(8) Test for mean of a Poisson distribution

**Option 1**:  $H_0: X \sim Po(\lambda)$ 

Test statistic: X, with the critical value obtained from the cumulative Poisson table (or calculated manually).

[Note: As for the Binomial proportion, we are using a single value - effectively from an infinite number of trials, with an infinitesimal probability of success.]

**Option 2**:  $H_0: X \sim Po(\lambda)$  approx.  $\sim N(\lambda, \lambda)$ 

Test statistic: X, with the critical value obtained from the Normal table. X should be >10. [A continuity correction is not usually required.]

(9) Test for variance of a Normal distribution

# Option 1

$$H_0: X \sim N(\mu, \sigma^2) \Rightarrow \frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$

Thus the test statistic is  $X^2 = \frac{(n-1)S^2}{\sigma^2}$ , with the critical value obtained from the  $\chi^2$  table.

#### Option 2

$$H_0: X \sim N(\mu, \sigma^2) \Rightarrow \frac{S^2}{\sigma^2} \sim F_{n-1,\infty}$$

Thus the test statistic is  $\frac{S^2}{\sigma^2}$ , with the critical value obtained from the *F* table.

### (B) Paired samples

(10) Test for difference of two means from Normal distributions

Treat differences as a single sample, determining their mean and variance, and proceed as for (A)(1)-(4). [In practice, the paired samples are likely to be small, with the population variance unknown, so that a *t*-test will be needed.]

(11) Test for difference between two unknown distributions, where the paired samples are small

 $H_0$ : The two samples are from a common distribution

Apply either the Wilcoxon signed-rank test or the Sign test to the differences (as in (5) and (6)), with M = 0, according to whether the distribution can be assumed to be symmetrical or not.

#### (C) Two independent samples

(12) Test for given difference between means of Normal distributions, with known common variance

$$H_0: \overline{X} - \overline{Y} \sim N(\mu_1 - \mu_2, \frac{\sigma^2}{n_1} + \frac{\sigma^2}{n_2})$$
  
Test statistic:  $z = \frac{(\overline{x} - \overline{y}) - (\mu_1 - \mu_2)}{\sigma \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$ 

Note: Often  $\mu_1 - \mu_2 = 0$ 

Variations (one or more of the following):

(i) Unknown common variance, with large samples: use  $s^2$ (estimated from pooled data:  $s^2 = \frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}$ ; this is an unbiased estimator)

(ii) Unknown common variance, and small samples (from Normal distributions): apply *t*-test with  $s^2 = \frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}$ 

and  $n_1 + n_2 - 2$  degrees of freedom

(iii) Unknown distributions, and large samples, so that approximate Normal distributions can be assumed, by the Central limit theorem

(iv) Different variances: replace  $\frac{\sigma^2}{n_1} + \frac{\sigma^2}{n_2}$  with  $\frac{{\sigma_1}^2}{n_1} + \frac{{\sigma_2}^2}{n_2}$ 

( $\sigma_1^2$  and  $\sigma_2^2$  can be estimated by  $s_1^2$  and  $s_2^2$ , if the samples are large)

(13) Test for difference between two unknown distributions, where the (independent) samples are small

Apply the Mann-Whitney test (related to the Wilcoxon Rank Sum test [see below] - not to be confused with the Wilcoxon Signed Rank test, used earlier (especially as both methods involve summing ranks).

Procedure

- (i)  $H_0$ : The two samples are from a common distribution
- (ii) Suppose that the sample sizes are m & n, where  $m \le n$

(iii) Rank the items in both samples together, with a rank of 1 for the smallest value. For example, one sample (of size 5) may contain the ranks 2, 4, 7, 11, 13; whilst the other (of size 8) has the ranks 1, 3, 5, 6, 8, 9, 10, 12.

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(iv) The test statistic is  $U = T - \frac{1}{2}m(m+1)$ ,

where T is the sum of ranks for the smaller sample (or either if m = n)  $\left[\frac{1}{2}m(m + 1) = 1 + \dots + m$  is the smallest possible value for T, so that  $U \ge 0$ ]

[Note: The Wilcoxon Rank Sum test has as its test statistic the T above, so that the critical values are those of the Mann-Whitney test, with  $\frac{1}{2}m(m+1)$  added.]

(v) Reject  $H_0$  if U < the lower tail critical value from the Mann-Whitney table [ie if U is suspiciously small if  $H_0$  is assumed]

[Note: For larger m & n,  $U \sim \operatorname{approx} N(\frac{1}{2}mn, \frac{1}{12}mn(m + n + 1))$ ; a continuity correction should be applied.]

(14) Test for equality of variance of two Normal distributions (given two independent samples)

Test statistic:  $F = \frac{{s_1}^2}{{s_2}^2}$ , where  ${s_1}^2 > {s_2}^2$ 

Reject  $H_0$  if F > upper tail critical value of  $F_{n_1-1,n_2-1}$ 

Note: To test for a given ratio  $\sigma_1^2 / \sigma_2^2$  of variances, test statistic becomes  $F = \frac{s_1^2 / \sigma_1^2}{s_2^2 / \sigma_2^2}$ 

#### (D) More than two independent samples

(15) Kruskal-Wallis test

Procedure:

(i)  $H_0$ : The samples are from a common distribution

(ii) Let the sample sizes be  $n_i$ , where i = 1 to k (and there are k samples), and let  $N = \sum_{i=1}^k n_i$ 

(iii) Rank the items in all the samples together, with a rank of 1 for the smallest value. For example, one of the samples (of size 5) may contain the ranks 2, 4, 7, 11, 13

(iv) The test statistic is  $H = \left[\frac{12}{N(N+1)}\sum_{i=1}^{k} \frac{T_i^2}{n_i}\right] - 3(N+1)$ ,

where  $T_i$  is the sum of ranks for the *i*th sample

(v) Reject  $H_0$  if H > upper tail critical value of  $\chi^2_{k-1}$